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## How to read a paper

# Statistics for the non-statistician. II: “Significant” relations and their pitfalls

Trisha Greenhalgh

This is the fifth in a series of 10 articles introducing non-experts to finding medical articles and assessing their value

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This article continues the checklist of questions that will help you to appraise the statistical validity of a paper. The first of this pair of articles was published last week.<sup>1</sup>

### Correlation, regression, and causation

Has correlation been distinguished from regression, and has the correlation coefficient ( $r$  value) been calculated and interpreted correctly?

For many non-statisticians, the terms “correlation” and “regression” are synonymous, and refer vaguely to a mental image of a scatter graph with dots sprinkled messily along a diagonal line sprouting from the intercept of the axes. You would be right in assuming that if two things are not correlated, it will be meaningless to attempt a regression. But regression and correlation are both precise statistical terms which serve quite different functions.<sup>1</sup>

The  $r$  value (Pearson’s product-moment correlation coefficient) is among the most overused statistical instrument. Strictly speaking, the  $r$  value is not valid unless the following criteria are fulfilled:

- The data (or, more accurately, the population from which the data are drawn) should be normally distributed. If they are not, non-parametric tests of correlation should be used instead.<sup>1</sup>
- The two datasets should be independent (one should not automatically vary with the other). If they are not, a paired  $t$  test or other paired test should be used.
- Only a single pair of measurements should be made on each subject. If repeated measurements are made, analysis of variance should be used instead.<sup>2</sup>
- Every  $r$  value should be accompanied by a P value, which expresses how likely an association of this strength would be to have arisen by chance, or a confidence interval, which expresses the range within which the “true”  $r$  value is likely to lie.

Remember, too, that even if the  $r$  value is appropriate for a set of data, it does not tell you whether the relation, however strong, is causal (see below).



### Summary points

An association between two variables is likely to be causal if it is strong, consistent, specific, plausible, follows a logical time sequence, and shows a dose-response gradient

A P value of  $< 0.05$  means that this result would have arisen by chance on less than one occasion in 20

The confidence interval around a result in a clinical trial indicates the limits within which the “real” difference between the treatments is likely to lie, and hence the strength of the inference that can be drawn from the result

A statistically significant result may not be clinically significant. The results of intervention trials should be expressed in terms of the likely benefit an individual could expect (for example, the absolute risk reduction)

The term “regression” refers to a mathematical equation that allows one variable (the target variable) to be predicted from another (the independent variable). Regression, then, implies a direction of influence, although—as the next section will argue—it does not prove causality. In the case of multiple regression, a far more complex mathematical equation (which, thankfully, usually remains the secret of the computer that calculated it) allows the target variable to be predicted from two or more independent variables (often known as covariables).

The simplest regression equation, which you may remember from your schooldays, is  $y=a+bx$ , where  $y$  is the dependent variable (plotted on the vertical axis),  $x$  is the independent variable (plotted on the horizontal axis), and  $a$  is the  $y$  intercept. Not many biological variables can be predicted with such a simple equation. The weight of a group of people, for example, varies with their height, but not in a linear way. I am twice as tall as my son and three times his weight, but although I am four times as tall as my newborn nephew I am much more than six times his weight. Weight, in fact, probably varies more closely with the square of someone’s height than with height itself (so a quadratic rather than a linear regression would probably be more appropriate).

Of course, even when the height-weight data fed into a computer are sufficient for it to calculate the regression equation that best predicts a person’s weight from their height, your predictions would still be pretty poor since weight and height are not all that closely

PETER BROWN

correlated. There are other things that influence weight in addition to height, and we could, to illustrate the principle of multiple regression, enter data on age, sex, daily calorie intake, and physical activity into the computer and ask it how much each of these covariables contributes to the overall equation (or model).

The elementary principles described here, particularly the criteria for the  $r$  value given above, should help you to spot whether correlation and regression are being used correctly in the paper you are reading. A more detailed discussion on the subject can be found elsewhere.<sup>2,3</sup>

*Have assumptions been made about the nature and direction of causality?*

Remember the ecological fallacy: just because a town has a large number of unemployed people and a very high crime rate, it does not necessarily follow that the unemployed are committing the crimes. In other words, the presence of an association between A and B tells you nothing at all about either the presence or the direction of causality. To show that A has caused B (rather than B causing A, or A and B both being caused by C), you need more than a correlation coefficient. The box gives some criteria, originally developed by Sir Austin Bradford Hill, which should be met before assuming causality.<sup>4</sup>

## Probability and confidence

*Have "P values" been calculated and interpreted appropriately?*

One of the first values a student of statistics learns to calculate is the P value—that is, the probability that any particular outcome would have arisen by chance. Standard scientific practice, which is entirely arbitrary, usually deems a P value of less than 1 in 20 (expressed as  $P < 0.05$ , and equivalent to a betting odds of 20 to 1) as "statistically significant" and a P value of less than 1 in 100 ( $P < 0.01$ ) as "statistically highly significant."

By definition, then, one chance association in 20 (this must be around one major published result per journal issue) will seem to be significant when it is not, and one in 100 will seem highly significant when it is really what my children call a "fluke." Hence, if you must analyse multiple outcomes from your data set, you need to make a correction to try to allow for this (usually achieved by the Bonferroni method<sup>5,6</sup>).

A result in the statistically significant range ( $P < 0.05$  or  $P < 0.01$ , depending on what is chosen as

the cut off) suggests that the authors should reject the null hypothesis (the hypothesis that there is no real difference between two groups). But a P value in the non-significant range tells you that either there is no difference between the groups or that there were too few subjects to demonstrate such a difference if it existed—but it does not tell you which.

The P value has a further limitation. Guyatt and colleagues, in the first article of their "Basic Statistics for Clinicians" series on hypothesis testing using P values, conclude: "Why use a single cut off point [for statistical significance] when the choice of such point is arbitrary? Why make the question of whether a treatment is effective a dichotomy (a yes-no decision) when it would be more appropriate to view it as a continuum?"<sup>7</sup> For a better assessment of the strength of evidence, we need confidence intervals.

*Have confidence intervals been calculated, and do the authors' conclusions reflect them?*

A confidence interval, which a good statistician can calculate on the result of just about any statistical test (the  $t$  test, the  $r$  value, the absolute risk reduction, the number needed to treat, and the sensitivity, specificity, and other key features of a diagnostic test), allows you to estimate for both "positive" trials (those that show a statistically significant difference between two arms of the trial) and "negative" ones (those that seem to show no difference), whether the strength of the evidence is strong or weak, and whether the study is definitive (obviates the need for further similar studies). The calculation and interpretation of confidence intervals have been covered elsewhere.<sup>8</sup>

If you repeated the same clinical trial hundreds of times, you would not get exactly the same result each time. But, on average, you would establish a particular level of difference (or lack of difference) between the two arms of the trial. In 90% of the trials the difference between two arms would lie within certain broad limits, and in 95% of the trials it would lie between certain, even broader, limits.

Now, if (as is usually the case) you conducted only one trial, how do you know how close the result is to the "real" difference between the groups? The answer is you don't. But by calculating, say, the 95% confidence interval around your result, you will be able to say that there is a 95% chance that the "real" difference lies between these two limits. The sentence to look for in a paper should read something like: "In a trial of the treatment of heart failure, 33% of the patients randomised to ACE inhibitors died, whereas 38% of those randomised to hydralazine and nitrates died. The point estimate of the difference between the groups [the best single estimate of the benefit in lives saved from the use of an ACE inhibitor] is 5%. The 95% confidence interval around this difference is -1.2% to 12%."

More likely, the results would be expressed in the following shorthand: "The ACE inhibitor group had a 5% (95% CI -1.2% to 12%) higher survival."

In this particular example, the 95% confidence interval overlaps zero difference and, if we were expressing the result as a dichotomy (that is, is the hypothesis "proved" or "disproved?") we would classify it as a negative trial. Yet as Guyatt and colleagues argue, there probably is a real difference, and it probably lies closer to 5% than either -1.2% or 12%. A more useful

### Tests for causation<sup>4</sup>

- Is there evidence from true experiments in humans?
- Is the association strong?
- Is the association consistent from study to study?
- Is the temporal relation appropriate (did the postulated cause precede the postulated effect)?
- Is there a dose-response gradient (does more of the postulated effect follow more of the postulated cause)?
- Does the association make epidemiological sense?
- Does the association make biological sense?
- Is the association specific?
- Is the association analogous to a previously proved causal association?

## Calculating the “bottom line” effects on an intervention

Group	Outcome event		Total
	Yes	No	
Control group	a	b	a + b
Experimental group	c	d	c + d

Control event rate (CER) = risk of outcome event in control group =  $a/(a + b)$   
 Experimental event rate (EER) = risk of outcome event in experimental group =  $c/(c + d)$   
 Relative risk reduction (RRR) =  $(CER - EER)/CER$   
 Absolute risk reduction (ARR) =  $CER - EER$   
 Number needed to treat (NNT) =  $1/ARR = 1/(CER - EER)$   
 Odds ratio = 
$$\frac{(\text{odds of outcome event } \vee \text{ odds of no event) in intervention group}}{(\text{odds of outcome event } \vee \text{ odds of no event) in control group}}$$

conclusion from these results is that “all else being equal, an ACE inhibitor is the appropriate choice for patients with heart failure, but the strength of that inference is weak.”<sup>9</sup>

Note that the larger the trial (or the larger the pooled results of several trials), the narrower the confidence interval—and, therefore, the more likely the result is to be definitive.

In interpreting “negative” trials, one important thing you need to know is whether a much larger trial would be likely to show a significant benefit. To determine this, look at the upper 95% confidence limit of the result. There is only one chance in 40 (that is, a 2½% chance, since the other 2½% of extreme results will lie below the lower 95% confidence limit) that the real result will be this much or more. Now ask yourself, “Would this level of difference be clinically important?” If not, you can classify the trial as not only negative but also definitive. If, on the other hand, the upper 95% confidence limit represented a clinically important level of difference between the groups, the trial may be negative but it is also non-definitive.

The use of confidence intervals is still relatively uncommon in medical papers. In one survey of 100 articles from three of North America’s top journals (the *New England Journal of Medicine*, *Annals of Internal Medicine*, and the *Canadian Medical Association Journal*), only 43 reported any confidence intervals, whereas 66 gave a P value.<sup>7</sup> An even smaller proportion of articles interpret their confidence intervals correctly. You should check carefully in the discussion section to see whether the authors have correctly concluded not only whether and to what extent their trial supported their hypothesis, but also whether any further studies need to be done.

## The bottom line

*Have the authors expressed the effects of an intervention in terms of the likely benefit or harm which an individual patient can expect?*

**Table 1** Bottom line effects: treatment and outcome<sup>10</sup>

Treatment	Outcome at 10 years	
	Dead	Alive
Medical treatment (n=1325)	404	921
Coronary artery bypass grafting (n=1324)	350	974

It is all very well to say that a particular intervention produces a “statistically significant difference” in outcome, but if I were being asked to take a new medicine I would want to know how much better my chances would be (in terms of any particular outcome) than they would be if I didn’t take it. Four simple calculations (if you can add, subtract, multiply, and divide you will be able to follow this section) will enable you to answer this question objectively and in a way that means something to the non-statistician. These calculations are the relative risk reduction, the absolute risk reduction, the number needed to treat, and the odds ratio.

To illustrate these concepts, and to persuade you that you need to know about them, consider a survey which Tom Fahey and his colleagues conducted recently.<sup>10</sup> They wrote to 182 board members of district health authorities in England (all of whom would be in some way responsible for making important health service decisions), asking them which of four different rehabilitation programmes for heart attack victims they would prefer to fund:

Programme A reduced the rate of deaths by 20%;

Programme B produced an absolute reduction in deaths of 3%;

Programme C increased patients’ survival rate from 84% to 87%;

Programme D meant that 31 people needed to enter the programme to avoid one death.

Of the 140 board members who responded, only three spotted that all four “programmes” in fact related to the same set of results. The other 137 preferred one or other of the programmes, thus revealing (as well as their own ignorance) the need for better basic training in epidemiology for health authority board members.

Let us continue with the example shown in table 1, which Fahey and colleagues reproduced from a study by Salim Yusuf and colleagues.<sup>11</sup> I have expressed the figures as a two by two table giving details of which treatment the patients received in their randomised trial and whether they were dead or alive 10 years later.

Simple mathematics tells you that patients receiving medical treatment have a chance of  $404/1324 = 0.305$  or 30.5% of being dead at 10 years. Let us call this risk  $x$ . Patients randomised to coronary artery bypass grafting have a chance of  $350/1325 = 0.264$  or 26.4% of being dead at 10 years. Let us call this risk  $y$ .

The relative risk of death—that is, the risk in surgically treated patients compared with medically treated controls—is  $y/x$  or  $0.264/0.305 = 0.87$  (87%).

The relative risk reduction—that is, the amount by which the risk of death is reduced by the surgery—is  $100\% - 87\% (1 - y/x) = 13\%$ .

The absolute risk reduction (or risk difference)—that is, the absolute amount by which surgical treatment reduces the risk of death at 10 years—is  $30.5\% - 26.4\% = 4.1\%$  (0.041).

The number needed to treat—how many patients need coronary artery bypass grafting in order to prevent, on average, one death after 10 years—is the reciprocal of the absolute risk reduction:  $1/ARR = 1/0.041 = 24$ .

Yet another way of expressing the effect of treatment is the odds ratio. Look back at the two by two table and you will see that the “odds” of dying compared with the odds of surviving for patients in the medical treatment group is  $404/921 = 0.44$ , and for

patients in the surgical group is  $350/974 = 0.36$ . The ratio of these odds will be  $0.36/0.44 = 0.82$ .

The general formulas for calculating these "bottom line" effects of an intervention, taken from Sackett and colleagues' latest book,<sup>12</sup> are shown in the box.

The outcome event can be desirable (cure, for example) or undesirable (an adverse drug reaction). In the latter case, it is semantically preferable to refer to numbers needed to harm and the relative or absolute increase in risk.

## Summary

It is possible to be seriously misled by taking the statistical competence (and/or the intellectual honesty) of authors for granted. Some common errors committed (deliberately or inadvertently) by the authors of papers are given in the final box.

The articles in this series are excerpts from *How to read a paper: the basics of evidence based medicine*. The book includes chapters on searching the literature and implementing evidence based findings. It can be ordered from the BMJ Bookshop: tel 0171 383 6185/6245; fax 0171 383 6662. Price £13.95 UK members, £14.95 non-members.

I am grateful to Mr John Dobby for educating me on statistics and for repeatedly checking and amending this article. Responsibility for any errors is mine alone.

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## Ten ways to cheat on statistical tests when writing up results

- Throw all your data into a computer and report as significant any relation where  $P < 0.05$
- If baseline differences between the groups favour the intervention group, remember not to adjust for them
- Do not test your data to see if they are normally distributed. If you do, you might get stuck with non-parametric tests, which aren't as much fun
- Ignore all withdrawals (drop outs) and non-responders, so the analysis only concerns subjects who fully complied with treatment
- Always assume that you can plot one set of data against another and calculate an "r value" (Pearson correlation coefficient), and assume that a "significant" r value proves causation
- If outliers (points which lie a long way from the others on your graph) are messing up your calculations, just rub them out. But if outliers are helping your case, even if they seem to be spurious results, leave them in
- If the confidence intervals of your result overlap zero difference between the groups, leave them out of your report. Better still, mention them briefly in the text but don't draw them in on the graph—and ignore them when drawing your conclusions
- If the difference between two groups becomes significant four and a half months into a six month trial, stop the trial and start writing up. Alternatively, if at six months the results are "nearly significant," extend the trial for another three weeks
- If your results prove uninteresting, ask the computer to go back and see if any particular subgroups behaved differently. You might find that your intervention worked after all in Chinese women aged 52-61
- If analysing your data the way you plan to does not give the result you wanted, run the figures through a selection of other tests

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## An unfortunate mistake A heartbeat too many

Torsade de pointes, torsade de pointes. The phrase kept floating through the back of my mind as I stood on the tenth tee during the golf club's charity day two years ago in May, experiencing runs of fast heart beats which seemed to ebb and flow.

I had been used to tachycardias since my teenage years when a visit to the town's library had led me to suspect Wolf Parkinson White syndrome, later confirmed when I resisted my boss's desire to inject me with digoxin during a bioavailability study while working as a registrar. But this was different and why now had the phrase torsades de pointes come to haunt me? The golf continued and my awareness of the arrhythmias receded. The connection though was not to be made until the following week's clinic.

I dislike May anyway because of seasonal hay fever and, like many doctors I suspect, I have often self medicated. On this occasion, however, I had overdone it adding erythromycin for purulent sputum to the terfenadine. So pleased was I to have found a non-sedating antihistamine that I have recommended it to many others including my clinic staff nurse to whom I had

recently given a prescription for her young son who seemed to be suffering. Sensibly she had checked this out with her own general practitioner, a former senior house officer of mine, who had warned her about possible interactions with other drugs, including erythromycin. The nurse had been aware of my tendency to self prescribe and become quite alarmed when she recalled collecting a bottle of erythromycin for me two days before the golf and hence her relief when I turned up at the next week's clinic unscathed.

I have no electrocardiographic proof that I was experiencing torsades but strongly suspect that this was the diagnosis, and that I was recollecting subliminally a report of this arrhythmia induced by a combination of terfenadine and erythromycin. It is certainly listed in the *ABPI Compendium*, and in view of the recent publicity surrounding the subject my near miss and reprehensible prescribing seemed worth reporting. At least the golf day raised plenty of money for our local dialysis unit.

M J Weston, *consultant physician, Chelmsford*